A global fit determination of effective $\Delta m^2_{31}$ from baseline dependence of reactor $\bar{\nu}_e$ disappearance

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Recently, three reactor neutrino experiments, Daya Bay, Double Chooz and RENO have directly measured the neutrino mixing angle $\theta_{13}$. In this Letter, another important oscillation parameter, effective $\Delta m^2_{31}$ ($= \Delta m^2_{31}$) is measured usingbaseline dependence of the reactor neutrino disappearance. A global fit is applied to publicly available data and $\Delta m^2_{31}$ is $2.95^{+0.42}_{-0.31} \times 10^{-3}$ eV$^2$, $\sin^2 2\theta_{13} = 0.099^{+0.016}_{-0.012}$ are obtained by setting both parameters free. This result is complementary to $\Delta m^2_{31}$ to be measured by spectrum shape analysis. The measured $\Delta m^2_{31}$ is consistent with $\Delta m^2_{32}$ measured by $\nu_e$ disappearance in MINOS, T2K and atmospheric neutrino experiments within errors. The minimum $\chi^2$ is small, which means the results from the three reactor neutrino experiments are consistent with each other.

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1. Introduction

Neutrino oscillation is a phenomenon which is not accounted for by the Standard Model of elementary particles, which assumes neutrinos as massless. There are six parameters in standard three flavor neutrino oscillation [1]: three mixing angles between flavor eigenstates and mass eigenstates ($\theta_{12}, \theta_{13}$ and $\theta_{23}$), one CP violating imaginary phase ($\delta$), and two independent squared mass differences: $\Delta m^2_{21} \equiv m^2_2 - m^2_1$, where $m_i$ are neutrino masses ($m_1, m_2, m_3$) of the three mass eigenstates ($\nu_e, \nu_\mu, \nu_\tau$) which correspond to the largest component of ($\nu_e, \nu_\mu, \nu_\tau$), respectively. $\theta_{12}$ and $\Delta m^2_{21}$ have been measured by solar neutrino disappearance experiments ($\nu_e \rightarrow \nu_e$) and long baseline reactor neutrino disappearance experiments ($\nu_e \rightarrow \nu_x, \nu_x \rightarrow \nu_e$) [2–4] and atmospheric experiments. All these measurements are summarized in [1]. Here, $\Delta m^2$ is a weighted average of $\Delta m^2_{21}$ and $\Delta m^2_{32}$, called absolute $\Delta m^2$ as described in detail later in this section. Recently, finite $\theta_{13}$ was finally measured by short baseline reactor neutrino experiments ($\nu_x \rightarrow \nu_x$) [2–4] and long baseline accelerator experiments ($\nu_x \rightarrow \nu_e$) [5,6].

Another effective mass squared difference $\Delta m^2_{31}$ can be measured by energy spectrum distortion and baseline dependence of the reactor-$\bar{\nu}_e$ experiments. This Letter is to measure $\Delta m^2_{31}$ by baseline dependence of the reactor neutrino-$\bar{\nu}_e$ experiments.

In reactor-$\bar{\nu}_e$ experiments, usually the neutrino disappearance is analyzed by a two flavor neutrino oscillation formula:

$$P(\bar{\nu}_e \rightarrow \bar{\nu}_e) = 1 - \sin^2 2\theta_{13} \sin^2 \frac{\Delta m^2_{31}}{4E_{\nu}}L,$$

where $L$ is baseline which is $\sim 1$ km and $E_{\nu}$ is neutrino energy, which is around a few MeV. $\Delta m^2_{31}$ is a weighted average of the two mass square differences, $|\Delta m^2_{31}|$ and $|\Delta m^2_{32}|$ of the standard parametrization,

$$\Delta m^2_{31} = c^2_{12}|\Delta m^2_{31}| + s^2_{12}|\Delta m^2_{32}|,$$

with $c_{ij}$ and $s_{ij}$ representing $\cos\theta_{ij}$ and $\sin\theta_{ij}$, respectively [7]. In the analyses of reactor-$\bar{\nu}_e$ experiments so far published, $\sin^2 2\theta_{13}$ is extracted assuming $\Delta m^2_{31} = \Delta m^2_{32}$ which is measured by MINOS experiment [8]. $\Delta m^2_{32}$ can be expressed as:

$$\Delta m^2_{32} = (s^2_{12} + s_{13}t_{23} \sin 2\theta_{12} \cos \delta) |\Delta m^2_{31}| + (c^2_{12} - s_{13}t_{23} \sin 2\theta_{12} \cos \delta) |\Delta m^2_{32}|,$$

where $t_{ij} = \tan \theta_{ij}$ [7]. Since there is a relation

$$\Delta m^2_{31} = \Delta m^2_{32} + \Delta m^2_{21},$$

in the standard three neutrino flavor scheme, the difference between $\Delta m^2_{31}$ and $\Delta m^2_{32}$ is expressed as follows,
where the overall sign depends on the mass hierarchy, and the ±0.3 term comes from the ambiguity of $\cos \delta$. The difference is much smaller than the current precisions of measurements and can be treated practically equivalent. A precision better than 1% is necessary to distinguish the mass hierarchy. However, if $\Delta m^2_{31}$ and $\Delta m^2_{32}$ are separately measured and if they turn out to be significantly different, it means the standard three-flavor neutrino scheme is wrong. Thus it is important to measure $\Delta m^2_{31}$ independently from $\Delta m^2_{32}$ to test the standard three-flavor neutrino oscillation.

The $E$ dependence and $L$ dependence analyses to extract $\Delta m^2_{31}$ use independent information, namely energy distortion and normalization and thus are complementary. Some of the authors demonstrated $\Delta m^2_{31}$ measurement using $L$ dependence of deficit value of each reactor-$\nu_{13}$ experiment in 2012 [9,10]. In this Letter the analysis is significantly improved by applying a detailed global fit making use of the publicly available information of the three reactor neutrino experiments.

In next section we re-analyze the published data of each experiment and compare with the results written in the papers in order to demonstrate our analysis produces an identical result. Section 3 discusses about potential correlations between the experiments. In Section 4, most recent Double Chooz, Daya Bay and RENO results [2,11,4] are combined and $\Delta m^2_{31}$ is extracted. Finally, a summary of this study is presented in Section 5.

2. Reactor neutrino data

Details of each experiment and their data are presented in this section and they are re-analyzed by the authors in order to demonstrate that the analysis methods used in this work are consistent with the publications from the experimental groups. The $\chi^2$ used in this section will be used to form a global $\chi^2$ function in Section 4.

2.1. Daya Bay

The Daya Bay (DB) reactor neutrino experiment consists of three experimental halls (EH), containing one or more antineutrino detectors (AD). The AD array sees 6 reactors clustered into 3 pairs: Daya Bay (DB1, DB2), Ling Ao (L1, L2) and Ling Ao-II (L3, L4) power stations. Fig. 1 shows the relative locations of reactors and AD and detector. All reactors are functionally identical pressurized water reactors with maximum thermal power of 2.9 GW [3].

In DB publication, the $\chi^2$ is defined as

$$\chi^2_{DB}(\theta_{13}, \Delta m^2_{31}) = \sum_d \frac{6}{M_d + B_d} \left[ M_d + B_d (1 + a + \sum_r \frac{6}{\sigma_r^2} a_r + \epsilon_d) \right]^2$$

$$+ \sum_r \frac{\sigma_r^2}{\sigma_r'^2} + \sum_d \left( \frac{\epsilon_d}{\sigma_d^2} + \frac{\eta_d^2}{(\sigma_d'^2)} \right),$$

where $M_d$ are the measured neutrino candidate events of the $d$-th AD with background subtracted, $B_d$ is the corresponding background, $T_d$ is the prediction from neutrino flux, Monte Carlo simulation (MC) and neutrino oscillation. $\sigma_d^2$ is the fraction of neutrino event contribution of the $r$-th reactor to the $d$-th AD determined by baselines and reactor fluxes. The uncorrelated reactor uncertainty is $\sigma_r$. $\sigma_d'$ is the uncorrelated detection uncertainty, and $\sigma_d'^2$ is the background uncertainty, with the corresponding pull-terms $(\sigma_r, \epsilon_d, \eta_d)$. An absolute normalization factor $a$ is determined from the fit to the data.

The values of $\sigma_d^2$ are not shown in Daya Bay publications and was estimated using

$$\sigma_d^2 = \frac{p_r / L_{rd}^2}{\sum_r (p_r / L_{rd}^2)} \quad \text{with} \quad p_r = \frac{w_r}{\sum_r w_r},$$

where $w_r$ is the thermal power of each reactor and $L_{rd}$ is the baseline of $r$-th reactor to $d$-th detector. In this analysis, the value of $p_r$ is considered 1/6 since all reactors have same nominal thermal power. The calculated $\sigma_d^2$ is shown in Table 2. All the others terms are shown in Table 3.

By using Eq. (6) and the data from Tables 2 and 3, we were able to reproduce Daya Bay’s result, where $T_d$ was multiplied by the value of the deficit probability ($P_{dr}$), defined as:

$$P_{dr}^{def} = 1 - \sin^2 2\theta \int_{0}^{\frac{8.0 \text{MeV}}{1.8 \text{MeV}}} \frac{1}{E} \frac{n_i(E) dE}{\int n_i(E) dE},$$

with $\Delta m^2$ being measured in eV$^2$, $L_{rd}$ in meters and $E$ in MeV. $n_i(E)$ is the expected energy spectrum of the observed neutrinos.
which is calculated by

\[ n_i(E) = S(E) \sigma_{\text{IBD}}(E). \]

(9)

where \( S_i(E_i) \) is the reactor neutrino spectrum per fission from fissile element \( i \) and \( \beta_i \) is a fraction of the fission rate of fissile element \( i \). For equilibrium light water reactors, \( \beta_i \) are similar and we use the values of Bugeney paper [12], namely \( {}^{235}\text{U} : 236\text{U} : 238\text{U} : 241\text{Pu} = 0.538 : 0.078 : 0.328 : 0.056 \). In this study, \( S_i(E_i) \) is approximated as an exponential of a polynomial function which is defined in [13],

\[ S_i(E_i) \propto \exp \left[ \sum_{j=1}^{6} \alpha_j E_i^{(j-1)} \right]. \]

(10)

\( \sigma_{\text{IBD}} \) is the cross section of the inverse process of neutron \( \beta \)-decay (IBD), that can be precisely calculated from the neutron lifetime [14]. The energy dependence of the IBD cross section is

\[ \sigma_{\text{IBD}}(E_i) = \left[ E_i (\text{MeV}) - 1.29 \right] \sqrt{E_i^2 - 2.59 E_i + 1.4}. \]

(11)

\( \sin^2 2\theta_{13} \) is extracted by fixing \( \Delta m^2_{31} \) as the MINOS \( \Delta m^2_{31} = 2.32 \times 10^{-3} \text{ eV}^2 \) [8]. The \( \chi^2 \) distributions of the Daya Bay paper and our calculation are compared in Fig. 2. The Daya Bay central value and uncertainty is \( \sin^2 2\theta_{13}^{\text{DB}} = 0.089 \pm 0.011 \) while our analysis showed \( \sin^2 2\theta_{13} = 0.090^{+0.011}_{-0.010} \) in good agreement with the published value. We also verified how different values for the fission rates coefficients of Eq. (10) and different assumptions for Eq. (7), affect the final result. Dependence on the burn-up values is less than 0.001, as it was determined by replacing the burn-up assumption with that of the Chooz reactors at the beginning and end of the reactor cycle. Extreme assumptions on Eq. (7) (one or two reactors off for the whole data period, for example) had an effect of less than 0.002 on the central value, with no change on the sensitivity. Moreover, the good agreement between the \( \chi^2 \) distributions, shows that the assumptions are reasonable.

### 2.2. RENO

The Reactor Experiment for Neutrino Oscillation (RENO) is located in South Korea and has two identical detectors, one near (ND) and one far (FD) from an array of six commercial nuclear reactors, as shown in Fig. 3.

Together with the distances of each detector reactor pair, the contribution of each reactor flux to each detector for the period of their first analysis is available [15] and are summarized in Table 4. The RENO \( \chi^2 \) is defined as:

\[
\chi^2_{\text{REO}}(\theta_{13}, \Delta m^2_{31}) = \begin{array}{c}
\sum_{d} \left[ \frac{N_{\text{obs}}^d - b_d - (1 + n_d + \xi_d) \sum_{r} (1 + f_r) N_{\text{exp}}^{d,r}}{\sigma_d^2} \right]^2 \\
+ \sum_{d} \left( \frac{\left( \xi_d \sigma_d^2 \right)^2}{\sigma_d^2} + \frac{\left( b_d \sigma_d^2 \right)^2}{\sigma_d^2} \right) + \sum_{r} \frac{f_r^2}{\sigma_r^2}
\end{array}
\]

(12)

where \( N_{\text{obs}}^d \) is the number of observed IBD candidates in each detector after background subtraction and \( N_{\text{exp}}^{d,r} \) is the number of
expected neutrino events, including detection efficiency, neutrino oscillations and contribution from the $r$-th reactor to each detector determined from baseline distances and reactor fluxes. A global normalization $\eta$ is taken free and determined from the fit to the data. The uncorrelated reactor uncertainty is $\sigma_{r}$, $\sigma_{d}$ is the uncorrelated detection uncertainty, and $\sigma_{BG}$ is the background uncertainty, and the corresponding pull parameters are $(\xi_{r}, \xi_{d}, b_{d})$. The values of these variables are presented in the Table 5.

The expected number of events for both detectors are not present in the RENO paper, but the ratio between data and expectation is shown. This ratio and the quantities of Table 4 were used to calculate the expectation value ($N_{\text{exp}}$).

Using the data in the Table 5, Eq. (8) and the MINOS $\Delta m_{32}^2$, we obtained $\sin^2 2\theta_{13} = 0.111 \pm 0.024$ which is in good agreement with their published value of $\sin^2 2\theta_{13}\text{RENO} = 0.113 \pm 0.023$. The $\chi^2$ distributions are also very similar as shown in Fig. 4.

### 2.3. Double Chooz

The Double Chooz (DC) experiment uses the two Chooz B reactors with thermal power of 4.25 GWe each. Currently, the experiment is using only the far detector, since its near detector is not complete yet. The Bugey-4 measurement [12] is used as a reference of the absolute neutrino flux in the analysis, and the relative location of the far detector and reactors are shown in Fig. 5, where the distances from the detector to each reactor are 998.1 and 1114.6 meters [16].

The Double Chooz Collaboration published a rate plus shape analysis result [2]. An effect of the shape analysis in this case is an evaluation of main backgrounds of $^9\text{Li}$ and fast neutron from the energy spectrum beyond the reactor neutrino energy range. Since information of detailed energy spectrum, which is necessary to reproduce the analysis, are not publicly available, we do not consider here the shape analysis but restrict only to the rate analysis. After the second publication on the $\sin^2 2\theta_{13}$ measurement, the Double Chooz group published a result of the direct measurement of backgrounds by making use of 7.53 days reactor-OFF period [17]. We used these data in addition to the background evaluation inputs written in [2] to improve the background estimation instead of the energy spectrum analysis. The relative neutrino-flux uncertainty for reactor-OFF period is much larger than reactor-ON period. The dominant uncertainty comes from long-life isotopes whose abundance are not well known. It has negligible contribution in reactor-ON period [17]. Therefore, we regard the error correlation on neutrino flux between the reactor ON and OFF periods to be uncorrelated. We performed a similar $\chi^2$ analysis as Daya Bay and RENO cases, assuming that the detector and background related uncertainties of [17] and [2] are fully correlated.

$$
\chi^2_{DC}(\theta_{13}, \Delta m_{32}^2) = \frac{2}{i=1} \frac{\left[ N_{\text{obs}} - (N_{\text{exp}}(1 + \alpha_i + \varepsilon) + B_i(1 + b)) \right]^2}{N_{\text{exp}} + B_i} + \sum_{i} \left( \frac{\alpha_i}{\alpha_{i}} \right)^2 + \frac{\varepsilon^2}{\sigma_{\varepsilon}^2} + \frac{b^2}{\sigma_{b}^2},
$$

(13)

where $N_{\text{obs}}$ is the number of the observed neutrino event candidates. The subscript “i” represents reactor-ON and OFF period. $N_{\text{exp}}$ is the number of expected neutrino events, including detection efficiency and oscillation effects, and $B$ is the total expected number of background events. The $\sigma_{r}$, $\sigma_{d}$, and $\sigma_{BG}$ are the reactor, detection and background uncertainties, respectively. The corresponding pull parameters are $(\alpha_i, \varepsilon, b)$. Using the parameters shown in Table 6, we obtained $\sin^2 2\theta_{13} = 0.131 \pm 0.048$ which is consistent with the result of the DC publication, $\sin^2 2\theta_{13} = 0.109 \pm 0.039$, although the background evaluation methods are different using different data sets. We also did a rate only analysis of the Double Chooz data, which result agreed with the published one.

### 3. Correlation evaluation of systematic uncertainties

In reactor neutrino experiments, the expected number of observed events ($N_{\text{exp}}$) is defined by:

$$
N_{\text{exp}} = \frac{1}{4\pi L^2} N_{\mu} \frac{P_{\text{eff}}}{\langle E_f \rangle} (\sigma_f),
$$

(14)

where $L$ is the reactor-detector baseline, $N_{\mu}$ is the number of targets in the detector, $\varepsilon$ is the detector efficiency, $P_{\text{eff}}$ is the reactor

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**Table 5**

| RENO: Fitting parameters. Differently from [4], here the BKGs are summed into a single quantity. The total BKG is subtracted from the IBD candidates giving $N_{\text{obs}}$, $N_{\text{exp}}$ is calculated as described in Section 2.2. |
|------------------|------------------|
| **ND** | **FD** |
| IBD candidates | 154.088 | 17.102 |
| $N_{\text{exp}}$ | 151723.54 | 17565.72 |
| Total BKG [day$^{-1}$] | 21.75 ± 5.93 | 4.24 ± 0.75 |
| Live Time [days] | 192.42 | 222.06 |
| Efficiency | 0.647 | 0.745 |
| $N_{\text{obs}}$ | 149.902.86 | 16.160.46 |
| $\sigma_d$ | 0.002 | 0.002 |
| $\sigma_e$ | 0.009 | 0.009 |
| $\sigma_B$ | 1114.05 | 166.54 |

**Table 6**

Double Chooz: Fitting parameters [2,17]. The detector uncertainty is the combination of detector response and efficiency uncertainties, and the BKGs are combined in a single quantity for each data set.

<table>
<thead>
<tr>
<th>Reactor-ON</th>
<th>Reactor-OFF</th>
</tr>
</thead>
<tbody>
<tr>
<td>IBD candidates</td>
<td>8449</td>
</tr>
<tr>
<td>IBD prediction</td>
<td>8439.6</td>
</tr>
<tr>
<td>Total BKG [day$^{-1}$]</td>
<td>$2.18 \pm 0.58$</td>
</tr>
<tr>
<td>Live Time [days]</td>
<td>227.93</td>
</tr>
<tr>
<td>$\sigma_d$</td>
<td>0.010</td>
</tr>
<tr>
<td>$\sigma_e$</td>
<td>0.017</td>
</tr>
</tbody>
</table>

**Fig. 5.** Relative locations of detector and reactors of Double Chooz experiment. Scale is approximate.
thermal power, \(\langle E_f \rangle\) is the mean energy released per fission, and \(\langle \sigma_f \rangle\) is the cross section per fission defined as:
\[
\langle \sigma_f \rangle = \sum_i \mu_i \int S_i(E) \sigma_{\text{BD}}(E) dE.
\]

For each experiment, \(L\), \(N_{\text{BD}}, \varepsilon\), and \(P_{\text{BD}}\) terms are determined independently. Therefore they can be assumed to be uncorrelated. On the other hand, \(\langle E_f \rangle\) and \(\langle \sigma_f \rangle\) terms are taken from the same references and the uncertainties of these terms are correlated between the experiments. From the Bugey and Chooz experimental results, the total uncertainty on spectrum prediction is 2.7%, where a 2% correlation is expected between the experiments as treated in [18]. Fully correlated signal prediction uncertainties between experiments, which come from neutrino flux and detection efficiency, can be canceled by overall normalization factors used in the analyses of the Daya Bay and RENO. It allows us only to take into account remaining uncertainties between detectors or periods for each experiment. Daya Bay and RENO treat the remaining uncertainties as uncorrelated in their publications.

4. Combined analysis

As explained before, the main method of this work is to combine all the data of the current neutrino reactor experiments in a single \(\chi^2\) function. Then we look for the minimum \(\chi^2\) value, calculate the \(\Delta \chi^2\) distribution, and determine the confidence level regions. The \(\chi^2\) function used for such analysis was chosen so as to use the data from Tables 1 to 6 as well as the correlation as described in Section 3. The definition of our global \(\chi^2\) is
\[
\chi^2_G = \chi^2_{\text{DB}} + \chi^2_{\text{RE}} + \chi^2_{\text{IBD}}.
\]
with the \(\chi^2\) of each experiment defined as in Section 2. Therefore, this function has 32 pull terms: 18 for Daya Bay (6 reactors, 6 detectors and 6 backgrounds), 10 for RENO (6 reactors, 2 detectors and 2 backgrounds) and 4 for Double Chooz (2 reactors, 1 detector and 1 background), which also contains the two overall normalization factors, one for Daya Bay and the other for RENO data set.

For all combinations of \(\Delta m^2\) and \(\sin^2 2 \theta\), the \(\chi^2_G\) is minimized with respect to the pull terms. Fig. 6 shows a map of the absolute \(\chi^2\) and Fig. 7 shows the \(\Delta \chi^2\) contour map near the \(\chi^2_{\text{min}}\), obtained by such procedure. From the minimum point and the 1 \(\sigma\) error region in the 1D \(\chi^2\) distribution,
\[
\chi^2_{\text{min}} = 5.14 / 6 \quad \text{Degrees of freedom},
\]
\[
\Delta m^2_{31} = 2.95^{+0.42}_{-0.61} \times 10^{-3} \text{ eV}^2,
\]
\[
\sin^2 2 \theta_{13} = 0.099^{+0.016}_{-0.012}.
\]
are obtained. This \(\Delta m^2_{31}\) is consistent with \(\Delta m^2_{32}\) measured by accelerator experiments [8,19], confirming the standard three flavor neutrino oscillation within the error. The \(\sin^2 2 \theta_{13}\) obtained here is independent from \(\Delta m^2_{32}\). The small \(\chi^2_{\text{min}}/\text{DoF}\) means the data from the three reactor neutrino experiments are consistent with each other.

All the pull terms output were within 1 \(\sigma\) from the input value, and the normalization factors obtained from the fit to the data, were both less than 1%.

In Fig. 8 the baseline dependence of the disappearance probability of each detector is shown, where the probability is calculated using the parameters output which give the best fit. The Double Chooz has a large effect on this \(\Delta m^2\) determination because it locates at a baseline where the slope of the oscillation is large. In the near future, when the near detector of the Double Chooz experiment starts operation, the accuracy of this \(\Delta m^2_{31}\) measurement is expected to improve much.

Complementary to this study, we demonstrated a similar, but simpler and robust measurement of the effective \(\Delta m^2_{31}\) from the baseline dependence of the disappearance probabilities of the three reactor-\(\theta_{13}\) experiments [9,10]. The result obtained on that work of \(\Delta m^2_{31} = 2.99^{+1.13}_{-1.58} \times 10^{-3} \text{ eV}^2\), is compatible with the value obtained in this Letter. In addition, a similar \(\Delta \chi^2\) distribution is presented in [20, Fig. 4]. However, the central value of \(\Delta m^2_{31}\) could not be compared since only the distribution is presented.

5. Summary

In this work, a global fit of the data from all the current reactor-\(\theta_{13}\) experiments was performed to measure \(\Delta m^2_{31}\). The combination of the data from Daya Bay, RENO and Double Chooz resulted in \(\Delta m^2_{31} = 2.95^{+0.42}_{-0.61} \times 10^{-3} \text{ eV}^2\). This is consistent with \(\Delta m^2_{32}\) and it confirms that the experiments are observing standard three flavor neutrino oscillations within the error. The mixing angle obtained in this analysis is \(\sin^2 2 \theta_{13} = 0.099^{+0.016}_{-0.012}\). The small \(\chi^2_{\text{min}}/\text{DoF}\) value indicates that the data from the three reactor experiments are consistent with each other. This analysis uses
Fig. 8. Reactor $\bar{\nu}_e$ survival probabilities. The solid line is the oscillation pattern obtained in this analysis and dot-dashed line uses MINOS $\Delta m^2_{31}$ and the $\sin^2 2\theta_{13}$ that returns the minimum $\chi^2$. The data points are below the $\Delta m^2_{31}$ because they are calculated using the parameters returned by the best fit solution. Generally, a detector sees several reactors. The horizontal axis is a weighted baseline $\langle L \rangle$ and the horizontal bar in each data point shows the standard deviation of the distribution of the baselines, which is defined by $\sigma_l = \sqrt{\langle L^2 \rangle - \langle L \rangle^2}$, where $\langle L^n \rangle = \sum_k P_k L_k^n / \sum_k P_k L_k^n$, $k$ is the reactor index and $I_k$ and $P_k$ are the baseline and thermal power of the reactor $k$.

independent information from the energy spectrum distortion and it is possible to improve the accuracy of $\Delta m^2_{31}$ combining with results from energy spectrum analysis. It will be important to perform this kind of analysis to improve $\Delta m^2_{31}$ accuracy and to check the consistency of the results from the reactor-$\theta_{13}$ experiments.

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